

# CANADA'S NEXT top model



## Mixed models

and why sociolinguists should use them

Daniel Ezra Johnson

VARBRUL / GoldVarb	other
dependent variable (DV)	DV, response, y
factor group, independent variable (IV)	IV, factor (categorical), predictor, x
factor	level
factor weight	coefficient, effect, estimate, $\beta$
factor weight range	similar to 'effect size'
input probability	intercept
applications / total	(response) proportion

lmer	other
mixed model	mixed-effects, hierarchical, or multilevel model
fixed effect	main effect
(all) fixed-effects model	flat model
conditional modes of random effects	random effect estimates, random effect BLUPs

## Terminological 'translations'

PROPERTIES OF DATA	GoldVarb	Rbrul	R	POSSIBLE ANALYSIS
response / DV: 2 categories	✓	✓	✓	logistic regression
response: 3+ categories			✓	ordinal, multinomial logistic
response: count			✓	Poisson regression, etc.
response: continuous		✓	✓	linear regression
predictor(s) / IV(s) : categorical	✓	✓	✓	(any)
predictor(s): continuous		✓	✓	(any)
predictor(s): have interactions	<i>hard</i>		✓	(any)
random intercept(s)	?	✓	✓	<b>mixed model</b>
random slope(s)	??		✓	mixed model
lots of data (need for speed)		✓	✓	
		<i>hard</i>	✓	plots and graphics
			✓	other statistical methods
	✓			“slash” operator
	?	?		user friendly

## Comparing Software Tools



GoldVarb



Rbrul



R

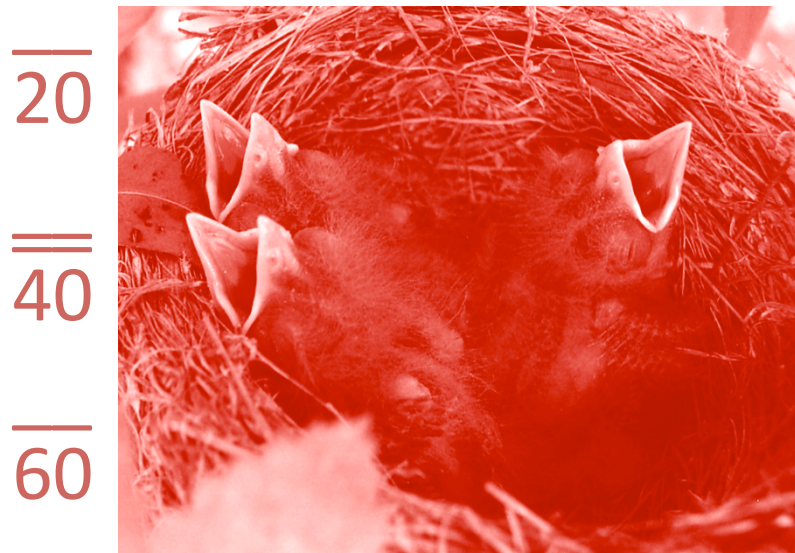
Finding the right tool for the job

- mixed models: both fixed effects and random effects
- fixed effect: ordinary regression predictor (IV)
- random effect: theoretically sampled from a population
  - est. population variance (s.d.) is the real parameter
  - individual estimates (BLUPs) “shrunk” towards mean
  - residual random effects should be normally distributed
- random intercept: individuals “high” or “low” (input prob.)
- random slope: individuals differ w.r.t. predictors (constraints)
- in model fitting, there is a penalty on the random effects
  - as much variance as possible assigned to fixed effects
  - only the left-over variance is assigned to random effects
- this random effect penalty allows nested models to fit
  - sometimes fixed vs. random (or separate runs) is a valid choice
  - but nested predictors must be random effects in a mixed model

## What are mixed models?



≠



?  
=



Mixed models for nested data

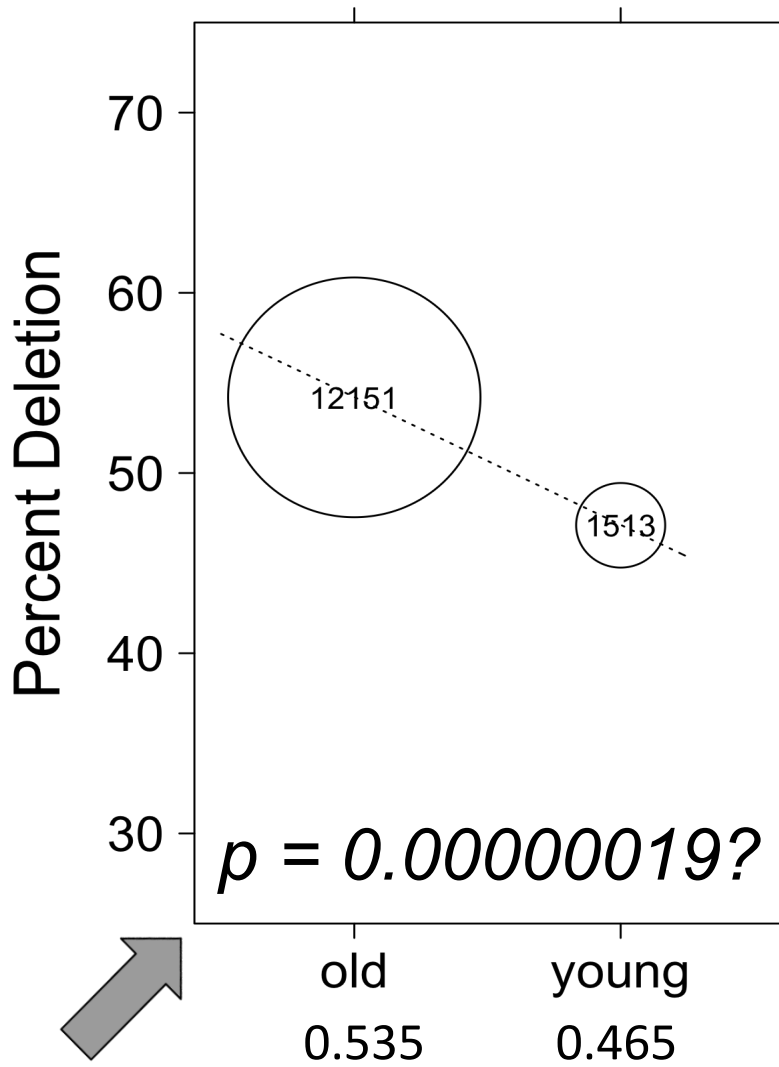


When we don't need mixed models

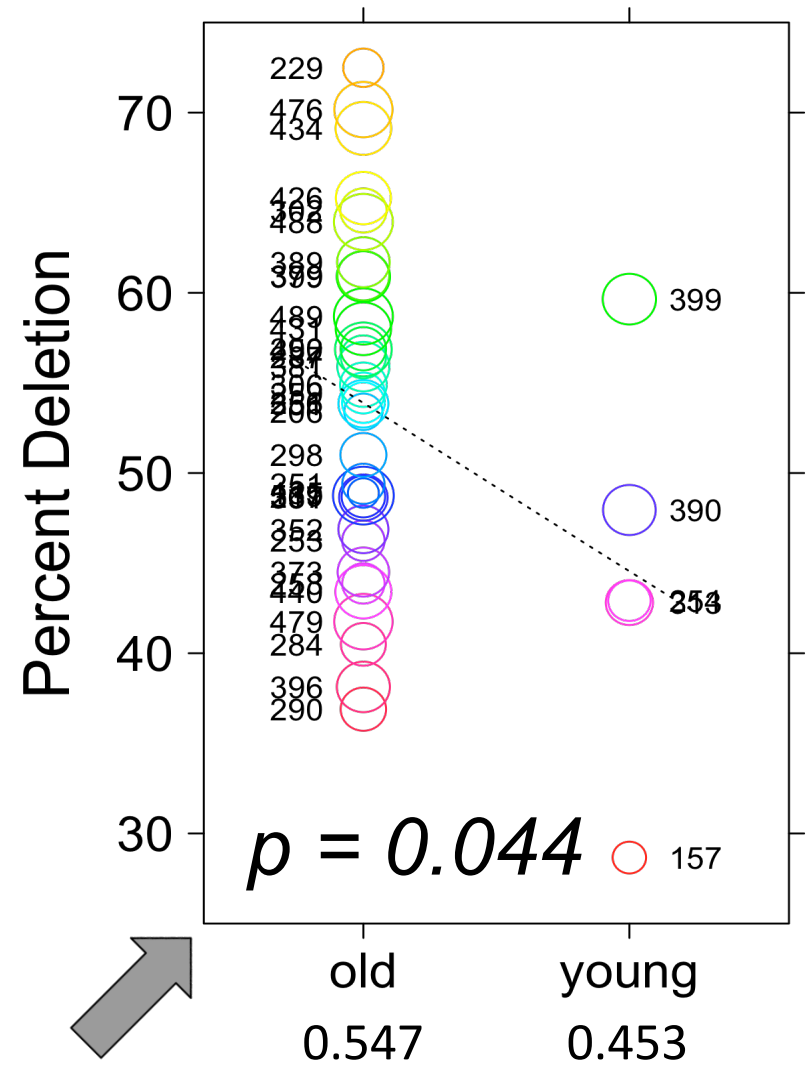


And when we might need them



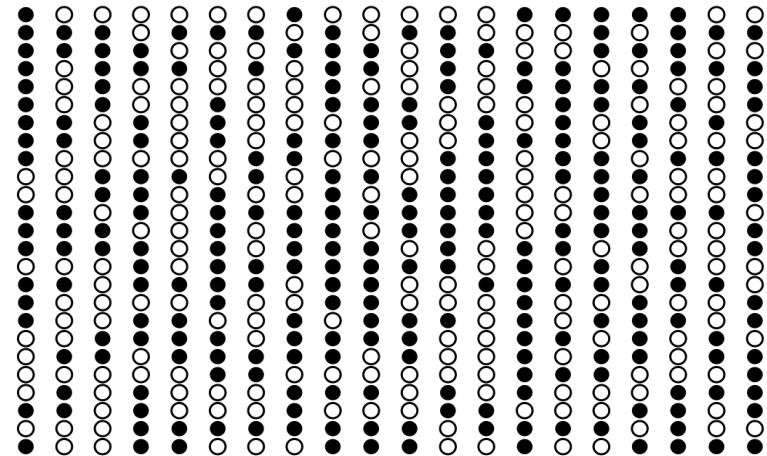
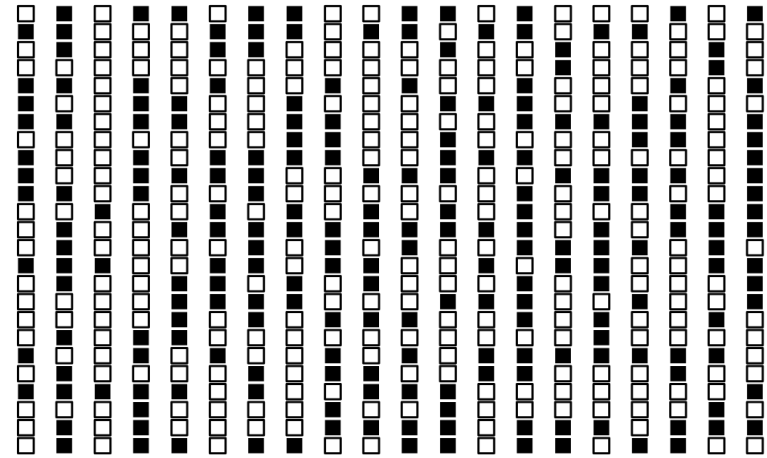


age w/ no random effect



age + random intercept: speaker

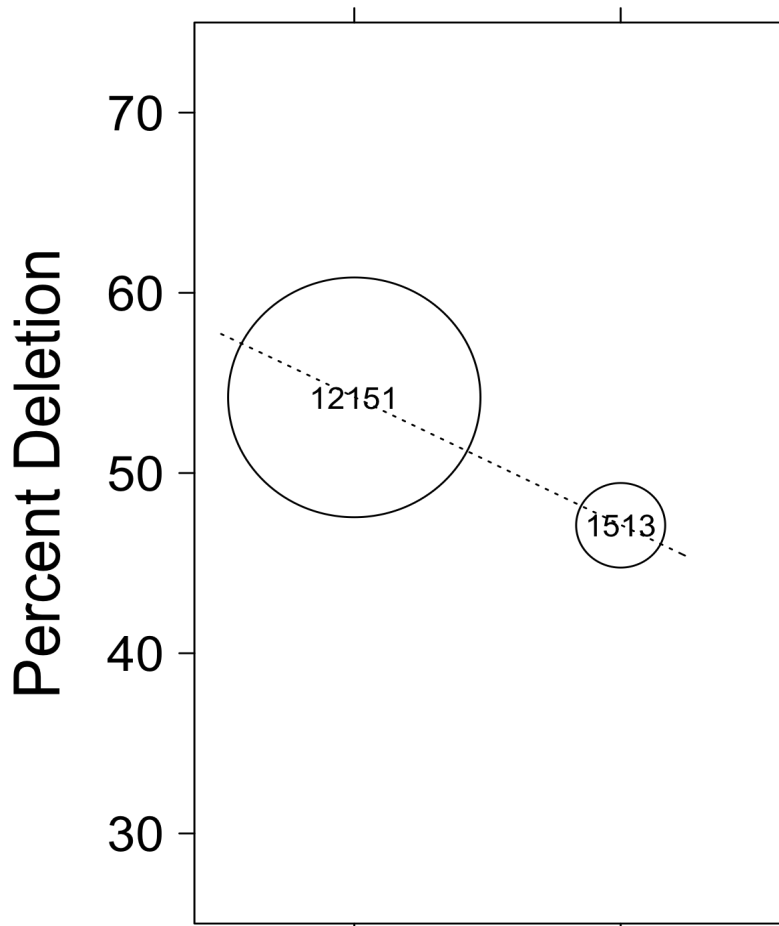
Random effects and significance



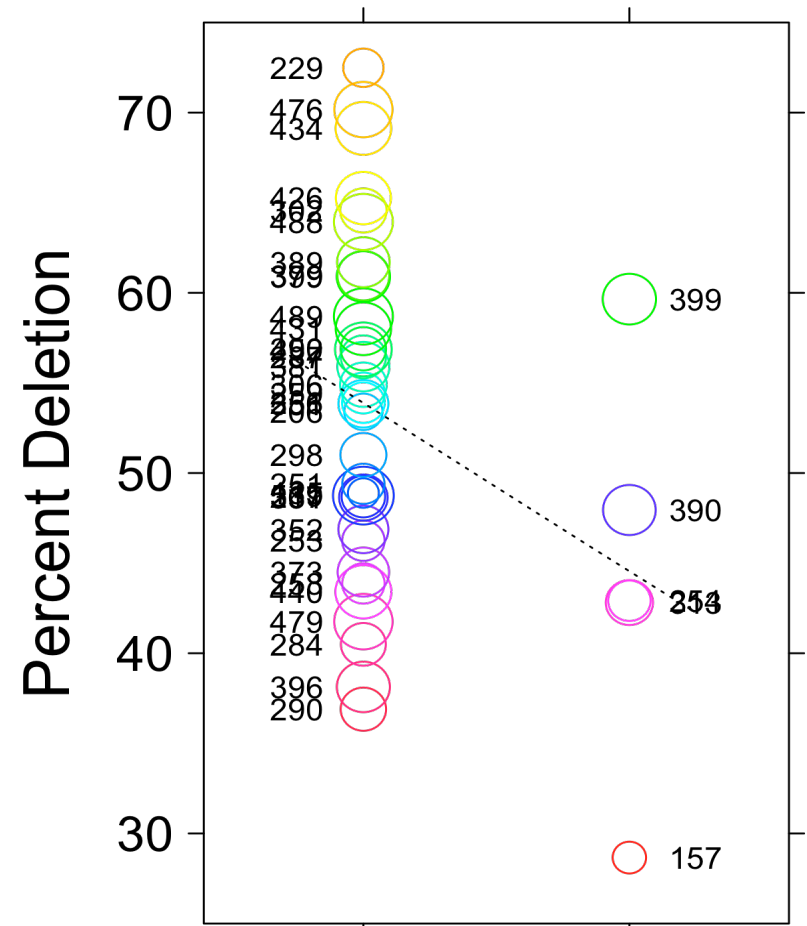
large effect size: 0.167 vs. 0.833  
small significance:  $p = 0.08$

small effect size: 0.45 vs. 0.55  
larger significance:  $p = 0.002$

# Significance vs. 'effect size'

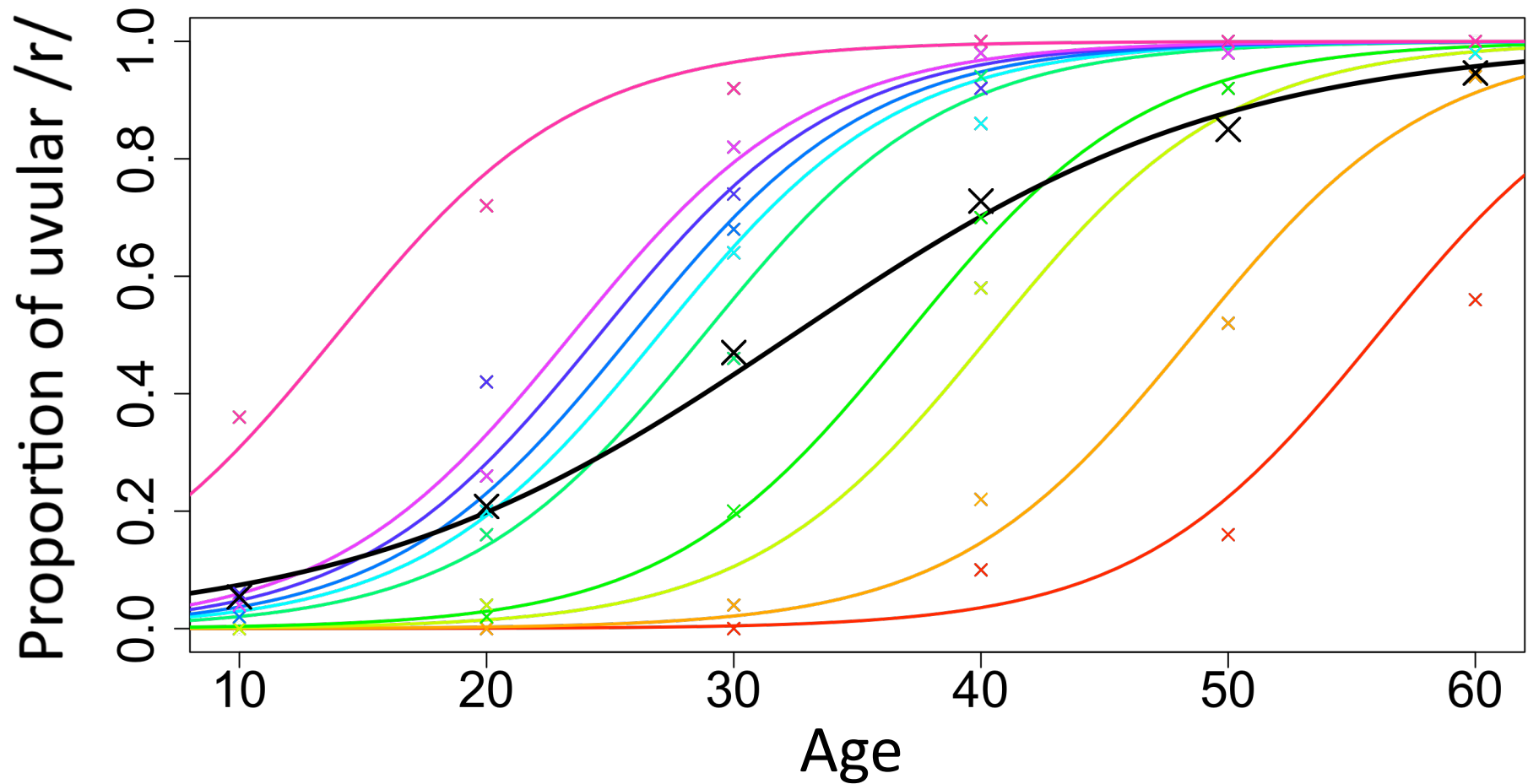


age w/ no random effect



age + random intercept: speaker

# Unbalanced data and effect size



age coefficient w/ no random effect: 0.113 log-odds/year  
 age coeff. w/ speaker random effect: 0.205 log-odds/year

# Crossed factors and effect size

## speaker-nesting predictors

*constant within (data from) each speaker*

age? gender race class c.o.p. ...

- significance more accurate:  
p = larger, “no longer significant”?
- effect sizes more accurate with unbalanced data: larger/smaller

## speaker-crossed predictors

*vary within (data from) each speaker*

age? style phon./gram. context...

- effect sizes more accurate:  
larger (logistic regression only)

# Summary: speaker effect's effects

## speaker-nesting predictors

## word-nesting predictors

*constant within (data from) each speaker*    *constant within (data from) each word*  
age? gender race class c.o.p. ...    frequency gram. cat. int. phon. ..

- significance more accurate:  
p = larger, “no longer significant”?
- effect sizes more accurate with  
unbalanced data, larger/smaller

## speaker-crossed predictors

## word-crossed predictors

*vary within (data from) each speaker*    *vary within (data from) each word*  
age? style phon./gram. context...    stress style ext. phon. ...

- effect sizes more accurate:  
larger (logistic regression only)

# Word effect just like speaker effect

## speaker-nesting predictors

*constant within (data from) each speaker*

age? gender race class c.o.p. ...

## word-nesting predictors

*constant within (data from) each word*

frequency gram. cat. int. phon. ...

- significance more accurate:  
p = larger, “no longer significant”?
- effect sizes more accurate with  
unbalanced data, larger/smaller

## speaker-crossed predictors

*vary within (data from) each speaker*

age? style phon./gram. context...

## word-crossed predictors

*vary within (data from) each word*

stress style ext. phon. ...

word



- effect sizes more accurate:  
larger (logistic regression only)

speaker



Crossed random effects for speaker & word

- use random effect estimates to identify ‘new’ fixed effects
  - modeled subject/word variation may include true individual variation, as well as unmodeled fixed effects
  -
- use random effect estimates to (empirically) build groups
- use random effect estimates as predictors in new models
- use random effect population variances to predict behavior of new subjects and words not in the original sample
- can perform an easy transformation into the ‘language’ of GoldVarb (with some caveats) – this is not a real problem

## Other benefits of mixed models



- cutting-edge statistics, like VARBRUL was in the 1970's
  - follow evolution on R-sig-ME
- double debate over p-values:
  - best way to calculate them
  - should they be used at all?
- convergence problems
  - requires more data (1000's > 100's)
- mixed model tool can be used well or badly, just like any model
  - still need to address multicollinearity
- should not be the only tool
  - mixed models are a better hammer, but everything is still not a nail
- “All models are wrong ... but some are useful.” – Box

# Drawbacks to mixed models

Substituting (2.16) into (2.15) into (2.6) provides the likelihood as

$$L(\beta, \theta, \sigma^2 | \mathbf{y}) = \prod_{i=1}^M \frac{\exp\left[-\|c_{0(i)} - R_{00(i)}\beta\|^2 / 2\sigma^2\right]}{(2\pi\sigma^2)^{n_i/2}} \text{abs}\left(\frac{|\Delta_i|}{|R_{11(i)}|}\right) \\ = \frac{\exp\left(-\sum_{i=1}^M \|c_{0(i)} - R_{00(i)}\beta\|^2 / 2\sigma^2\right)}{(2\pi\sigma^2)^{N/2}} \prod_{i=1}^M \text{abs}\left(\frac{|\Delta_i|}{|R_{11(i)}|}\right).$$

The term in the exponent has the form of a residual sum-of-squares for  $\beta$  pooled over all the groups. Forming another orthogonal-triangular decomposition

$$\begin{bmatrix} R_{00(1)} & c_{0(1)} \\ \vdots & \vdots \\ R_{00(M)} & c_{0(M)} \end{bmatrix} = Q_0 \begin{bmatrix} R_{00} & c_0 \\ \mathbf{0} & c_{-1} \end{bmatrix} \quad (2.17)$$

produces the reduced form

$$L(\beta, \theta, \sigma^2 | \mathbf{y}) \\ = (2\pi\sigma^2)^{-N/2} \exp\left(\frac{\|c_{-1}\|^2 + \|c_0 - R_{00}\beta\|^2}{-2\sigma^2}\right) \prod_{i=1}^M \text{abs}\left(\frac{|\Delta_i|}{|R_{11(i)}|}\right). \quad (2.18)$$

For a given  $\theta$ , the values of  $\beta$  and  $\sigma^2$  that maximize (2.18) are

$$\hat{\beta}(\theta) = R_{00}^{-1}c_0 \quad \text{and} \quad \hat{\sigma}^2(\theta) = \frac{\|c_{-1}\|^2}{N}, \quad (2.19)$$

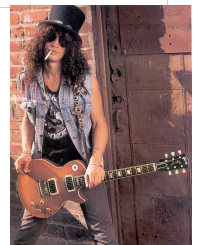
which give the profiled likelihood

$$L(\theta | \mathbf{y}) = L(\hat{\beta}(\theta), \theta, \hat{\sigma}^2(\theta) | \mathbf{y}) \\ = \left(\frac{N}{2\pi \|c_{-1}\|^2}\right)^{N/2} \exp\left(-\frac{N}{2}\right) \prod_{i=1}^M \text{abs}\left(\frac{|\Delta_i|}{|R_{11(i)}|}\right), \quad (2.20)$$

or the profiled log-likelihood

$$\ell(\theta | \mathbf{y}) = \log L(\theta | \mathbf{y}) \\ = \frac{N}{2} [\log N - \log(2\pi) - 1] - N \log \|c_{-1}\| + \sum_{i=1}^M \log \text{abs}\left(\frac{|\Delta_i|}{|R_{11(i)}|}\right). \quad (2.21)$$

The profiled log-likelihood (2.21) is maximized with respect to  $\theta$ , producing the maximum likelihood estimate  $\hat{\theta}$ . The maximum likelihood estimates  $\hat{\beta}$  and  $\hat{\sigma}^2$  are then obtained by setting  $\theta = \hat{\theta}$  in (2.19).



- it is fixed-effect models that make an assumption:
  - that residual subject and word variances are zero
  - i.e. that word-specific phonology is wrong
- mixed models are agnostic
  - random effects can be zero
  - they do not *assume* a word-specific (or speaker-specific) phonology, they *allow* for it *if it is supported by the data*
- must model speaker/word
  - with random effects, if nested
  - often crossed r. effects for both
- or other results will be wrong
  - maybe not very far wrong?
- as quantitative linguists, we strive for right numbers

Sali Tagliamonte  
 fellow panelists  
 Josef Fruehwald  
 Maryam Bakht  
 Meghan Armstrong  
 Kyle Gorman  
 Kirk Hazen  
 David Sankoff  
 Florian Jaeger  
 Rbrul testers  
 R developers



Doug Bates `lmer`  
 Qdoba on Bleecker

Pinheiro, José C. and Douglas M. Bates. 2000. *Mixed-Effects Models in S & S-PLUS*. New York: Springer.

Baayen, R. Harald, Douglas J. Davidson and Douglas M. Bates. 2008. Mixed-effects modeling with crossed random effects for subjects and items. *Journal of Memory and Language* 59, 390-412.  
 [I recommend this whole special issue on Emerging Data Analysis.]

Johnson, Daniel Ezra. 2009. Getting off the GoldVarb Standard: introducing Rbrul for mixed-effect variable rule analysis. *Language and Linguistics Compass* 3/1: 359-383.

Rbrul (a work in progress) is at:  
[www.danielezrajohnson.com/Rbrul.R](http://www.danielezrajohnson.com/Rbrul.R)

# Conclusions, thanks, references